

# Supplemental document: Fractional Gaussian Fields for Modeling and Rendering of Spatially-Correlated Media

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This document is supplemental to the paper entitled *Fractional Gaussian Fields for Modeling and Rendering of Spatially-Correlated Media*. In the following sections, we provide derivations of some important formulas and detailed discussions of some conclusions, as well as additional results.

## 1 DERIVATION OF $C(H, 1)$

The scaling term  $C(H, d)$  defined on  $\mathbb{R}^d$  is given by

$$C(H, d) = \frac{2^{-2H-d}\Gamma(-H)}{\pi^{d/2}\Gamma(H+d/2)}. \quad (1)$$

When  $d = 1$ , we have

$$\begin{aligned} C(H, 1) &= \frac{2^{-2H-1}\Gamma(-H)}{\pi^{1/2}\Gamma(H+1/2)} \\ &= \frac{2^{-2H-1}\Gamma(-H)\Gamma(H)}{\pi^{1/2}\Gamma(H+1/2)\Gamma(H)} \\ &= -\frac{2^{-2H-1}\frac{\pi}{H\sin(\pi H)}}{\pi^{1/2}2^{1-2H}\pi^{1/2}\Gamma(2H)} \\ &= -\frac{1}{4H\Gamma(2H)\sin(\pi H)} \\ &= -\frac{1}{2\Gamma(2H+1)\sin(\pi H)}. \end{aligned} \quad (2)$$

The third equation is based on the Euler's reflection formula

$$\Gamma(1-H)\Gamma(H) = \frac{\pi}{\sin(\pi H)} \quad (3)$$

and the duplication formula

$$\Gamma(H+1/2)\Gamma(H) = 2^{1-2H}\pi^{1/2}\Gamma(2H). \quad (4)$$

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Note that  $C(H, 1)$  satisfy

$$C(H+1, 1) = \frac{C(H, 1)}{(2H+1)(2H+2)}. \quad (5)$$

## 2 DERIVATION OF $\text{var}_p[\bar{\sigma}_t]$

Defining  $\mathbf{x}' = \mathbf{x} + t'\boldsymbol{\omega}$  and  $\mathbf{x}'' = \mathbf{x} + t''\boldsymbol{\omega}$ , the variance of  $\bar{\sigma}_t(\mathbf{x}) = [\int_0^t \sigma_t(\mathbf{x} + t'\boldsymbol{\omega})dt'] / t$  is calculated as

$$\text{var}[\bar{\sigma}_t] = \frac{1}{t^2} \int_0^t \int_0^t \text{cov}(\mathbf{x}', \mathbf{x}'')dt'dt''. \quad (6)$$

Substituting the autocovariance function of 1D pink noise into the above formula, we have

$$\begin{aligned} \text{var}_p[\bar{\sigma}_t] &= \frac{1}{t^2} \int_0^t \int_0^t C(H)S_w |t' - t''|^{2H} dt' dt'' \\ &= \frac{1}{t^2} C(H)S_w \int_0^t \int_0^t |t' - t''|^{2H} dt' dt'' \\ &= \frac{C(H)S_w}{(2H+1)(H+1)} t^{2H} \\ &= \frac{-S_w t^{2H}}{2(2H+1)(H+1)\Gamma(2H+1)\sin(\pi H)}. \end{aligned} \quad (7)$$

Using the fact that

$$2(2H+1)(H+1)\Gamma(2H+1) = \Gamma(2H+3) \quad (8)$$

we can simplify the above expression to

$$\begin{aligned} \text{var}_p[\bar{\sigma}_t] &= \frac{-S_w t^{2H}}{\Gamma(2H+3)\sin(\pi H)} \\ &= -2C(H+1)S_w t^{2H}. \end{aligned} \quad (9)$$

## 3 DERIVATION OF $\text{var}_f[\bar{\sigma}_t]$

For 1D fBm, we have

$$\begin{aligned} \text{var}_f[\bar{\sigma}_t(x)] &= \frac{1}{t^2} \int_0^t \int_0^t C(H)(|x' - x''|^{2H} - |x'|^{2H} - |x''|^{2H})dt'dt'' \\ &= \frac{2C(H)}{2H+1} S_w \left\{ \frac{t^{2H}}{2H+2} - \frac{(x+t)^{2H+1} - x^{2H+1}}{t} \right\} \end{aligned} \quad (10)$$

in which  $x' = x + t'$  and  $x'' = x + t''$ . Note that this expression depends on the spatial position  $x$  and has a never-ending growth with respect to  $x$  as we claimed in the paper.

To use fBm in practice, it is required to define fBm on a limited scale from 0 to an outer-scale  $L$ . Now, performing spatial averaging on the above expression, we get

$$\begin{aligned} \text{var}_f[\bar{\sigma}_t] &= \frac{1}{L} \int_0^L \frac{2C(H)}{2H+1} S_w \left\{ \frac{t^{2H}}{2H+2} - \frac{(x+t)^{2H+1} - x^{2H+1}}{t} \right\} \\ &= -\frac{2C(H)}{2H+1} S_w \frac{(L+t)^{2H+2} - L^{2H+2} - t^{2H+2} - t^{2H+1}L}{(2H+2)tL}. \end{aligned} \quad (11)$$

Using the expansion  $(L+t)^{2H+2} = L^{2H+2} + (2H+2)tL^{2H+1} + \dots + t^{2H+2}$ , we further arrive at

$$\begin{aligned} \text{var}_f[\bar{\sigma}_t] &= -\frac{2C(H)}{2H+1} S_w \frac{(2H+2)tL^{2H+1} + \dots - t^{2H+1}L}{(2H+2)tL} \\ &= -\frac{2C(H)}{2H+1} S_w \frac{(2H+2)L^{2H+1} + \dots - t^{2H}L}{(2H+2)L}. \end{aligned} \quad (12)$$

If we assume  $L \gg t$ , we will obtain

$$\text{var}_f[\bar{\sigma}_t] \approx -\frac{2C(H)}{2H+1} S_w L^{2H} \quad (13)$$

considering that terms in “...” all contain  $t$ . This expression is consistent with the one derived in the paper using the one-point scale-independence property of fBm.

#### 4 ONE-POINT SCALE-INDEPENDENCE OF FBM

In this section, we prove the one-point scale-independence property of fBm [Davis and Marshak 2004]. In the context of random extinction field  $\sigma_t$  of a fBm type, the property of one-point scale-independence requires:

- (1) The ensemble average of the line-averaged extinction is the same as the ensemble average of the extinction itself, i.e.,

$$\langle \bar{\sigma}_t \rangle = \langle \sigma_t \rangle. \quad (14)$$

- (2) The variance of the line-averaged field and the variance of the field itself differ at most by a small amount on the order of a very small ratio, i.e.,

$$\frac{\text{var}[\bar{\sigma}_t]}{\text{var}[\sigma_t]} - 1 = O\left(\left(\frac{t}{L}\right)^{2H}\right). \quad (15)$$

It is easy to prove the first requirement. For the second requirement, we know that the variance of the line-averaged field is given by

$$\begin{aligned} \text{var}_f[\bar{\sigma}_t] &= -\frac{2C(H)}{2H+1} S_w \frac{(L+t)^{2H+2} - L^{2H+2} - t^{2H+2} - t^{2H+1}L}{(2H+2)tL} \end{aligned} \quad (16)$$

while the variance of the field itself is

$$\begin{aligned} \text{var}_f[\sigma_t] &= \frac{1}{L} \int_0^L -2C(H) S_w |x|^{2H} dx \\ &= -\frac{2C(H) S_w}{2H+1} L^{2H}. \end{aligned} \quad (17)$$

Their ratio is

$$\begin{aligned} \frac{\text{var}[\bar{\sigma}_t]}{\text{var}[\sigma_t]} &= \frac{(L+t)^{2H+2} - L^{2H+2} - t^{2H+2} - t^{2H+1}L}{(2H+2)tL^{2H+1}} \\ &= \frac{1}{2H+2} \left\{ \frac{L}{t} \left[ \left(1 + \frac{t}{L}\right)^{2H+2} - 1 \right] \right. \\ &\quad \left. - \left(\frac{t}{L}\right)^{2H+1} - \left(\frac{t}{L}\right)^{2H} \right\}. \end{aligned} \quad (18)$$

When  $L \gg t$ ,  $\frac{t}{L} \rightarrow 0$ , we have

$$\lim_{\frac{t}{L} \rightarrow 0} \frac{L}{t} \left[ \left(1 + \frac{t}{L}\right)^{2H+2} - 1 \right] = 2H+2. \quad (19)$$

Then,

$$\begin{aligned} \frac{\text{var}[\bar{\sigma}_t]}{\text{var}[\sigma_t]} - 1 &= -\frac{1}{2H+2} \left[ \left(\frac{t}{L}\right)^{2H+1} + \left(\frac{t}{L}\right)^{2H} \right] \\ &= O\left(\left(\frac{t}{L}\right)^{2H}\right). \end{aligned} \quad (20)$$

#### 5 DERIVATION OF $\text{var}_{kf}[\bar{\sigma}_t]$

Based on the property of one-point scale-independence of  $k$ -fBm, we can obtain  $\text{var}_{kf}[\bar{\sigma}_t]$  using the variance of the field  $\sigma_t$  itself. In the case of  $k$ -fBm, the variance of  $\sigma_t$  is given by

$$\begin{aligned} \text{var}_{kf}[\sigma_t] &= \frac{1}{L} \int_0^L -2C(H) S_w \sum_{j=0}^{k-1} (-1)^j \binom{2H}{j} x^{2H} dx \\ &= -\frac{2C(H)}{2H+1} S_w \sum_{j=0}^{k-1} (-1)^j \binom{2H}{j} L^{2H} \\ &= \frac{2C(H) S_w (-1)^k}{2H+1} \binom{2H-1}{k-1} L^{2H} \end{aligned} \quad (21)$$

using the fact that

$$\sum_{j=0}^{k-1} (-1)^j \binom{2H}{j} = (-1)^{k-1} \binom{2H-1}{k-1}. \quad (22)$$

This is also the expression of  $\text{var}_{kf}[\bar{\sigma}_t]$ .

#### 6 ADDITIONAL RESULTS WITH ANISOTROPIC PHASE FUNCTIONS

In this section, we show the influence of phase functions on the appearance of spatially-correlated media. In Fig. 1, we choose the Henyey-Greenstein (HG) phase function parameterized by  $g$  (the asymmetry parameter) and render a homogeneous medium under different settings. Recall that negative values of  $g$  correspond to back-scattering while positive values correspond to forward-scattering. As seen, the phase function has a remarkable effect on the appearance when the  $H$  parameter is small. The influence weakens as  $H$  increases.

#### 7 DISCUSSION ON NEGATIVE CORRELATIONS

If we use the pink-noise type transmittance function in our model and simply let  $H \in (-1, -1/2)$ , we will get faster-than-exponential attenuations as shown in Fig. 2. This is a typical characteristic of

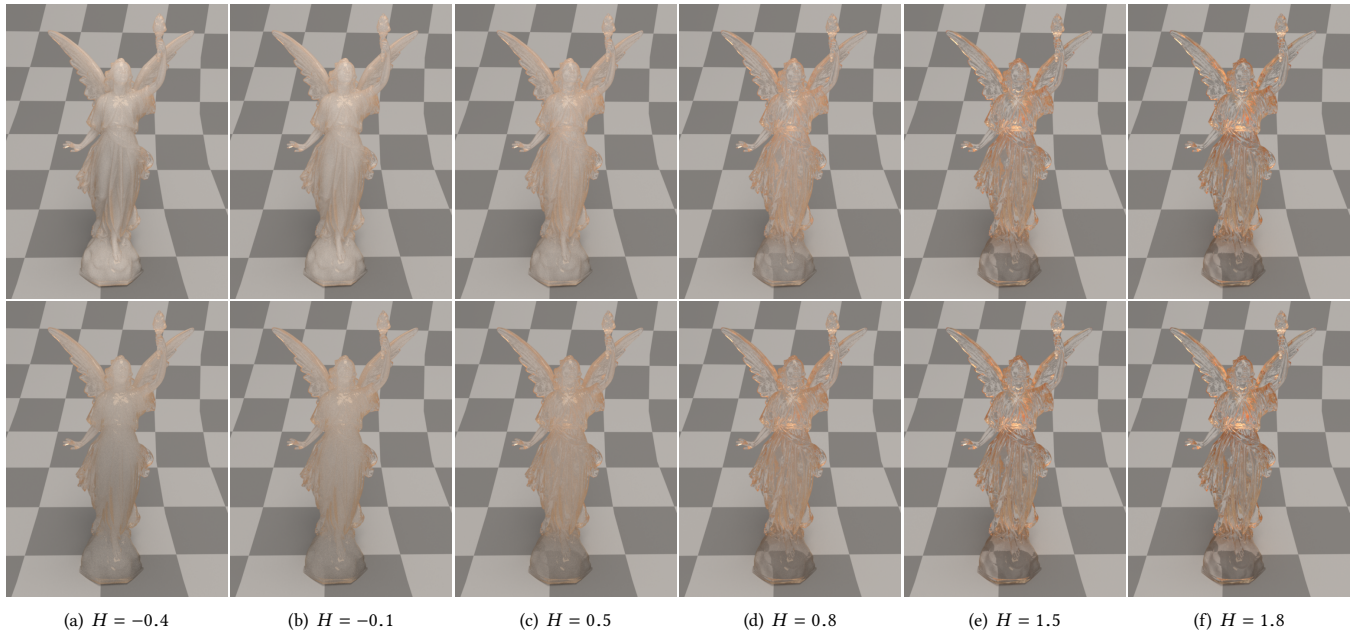


Fig. 1. Effects of spatial correlations in random media with anisotropic HG phase functions. The symmetry parameter  $g$  is set to  $-0.5$  (top row) and  $0.5$  (bottom row), respectively.

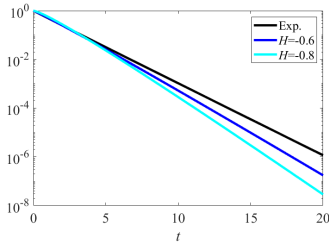


Fig. 2. Negative correlations can be achieved empirically by set  $H \in (-1, -1/2)$  in the pink-noise type transmittance function.



Fig. 3. Comparison between negative (left) and positive (right) correlations.

the negatively-correlated media, implying that our model can also support negative correlations to a certain extent. See the visual comparison in Fig. 3. However, we should emphasize that this simple extension is not physically-based because when  $H < -1/2$  we no longer have a straightforward Wiener-Khinchin connection between the PSD and the autocovariance function [Davis and Mineev-Weinstein 2011].

## 8 DISCUSSION ON WHITE NOISE

Our method converges to the classical transport with exponential falloff when the FGF is white noise. For white noise ( $H = -1/2$ ), the variance of  $\bar{\sigma}_t$  reduces to

$$\text{var}[\bar{\sigma}_t] = S_w t^{-1} \quad (23)$$

and the transmittance is given by

$$\text{Tr}(t) = \left(1 + \frac{S_w}{\sigma_m}\right)^{-\frac{\sigma_m^2}{S_w} t}. \quad (24)$$

Rearranging the above equation, we get

$$\text{Tr}(t) = e^{-\ln\left(1 + \frac{S_w}{\sigma_m}\right) \frac{\sigma_m^2}{S_w} t} \quad (25)$$

which means the transmittance is exponential in this case and the effective extinction is given by  $\ln\left(1 + \frac{S_w}{\sigma_m}\right) \frac{\sigma_m^2}{S_w}$ . Fig. 4 verifies that our model converges to the classical exponential transmittance with extinction  $\ln\left(1 + \frac{S_w}{\sigma_m}\right) \frac{\sigma_m^2}{S_w}$  when  $H$  approaches  $-1/2$ .

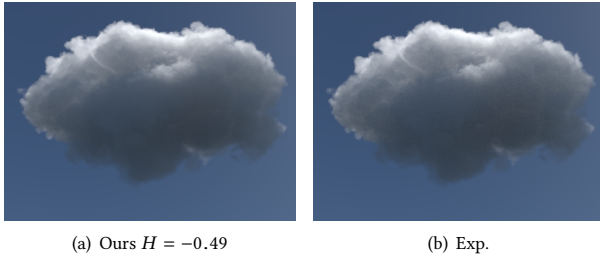


Fig. 4. Our method with a Hurst parameter close to  $-1/2$  (a) converges to the classical transport with exponential falloff (b).

If we further assume that  $S_w$  is very small (much smaller than  $\sigma_m$ ), the above expression of transmittance simplified to

$$\text{Tr}(t) = -e^{\sigma_m t} \quad (26)$$

based on the fact  $\ln\left(1 + \frac{S_w}{\sigma_m}\right) \approx \frac{S_w}{\sigma_m}$ . This gives the classical transmittance without micro-scale fluctuations.

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