

Total variation distance:

For distributions μ & π on finite \mathcal{X} ,

$$d_{TV}(\mu, \pi) = \frac{1}{2} \sum_{x \in \mathcal{X}} |\mu(x) - \pi(x)|$$

$$= \max_{S \subseteq \mathcal{X}} \mu(S) - \pi(S)$$

choose

$$S = \{x : \mu(x) > \pi(x)\}$$

$\rightarrow S \subseteq \mathcal{X}$

$$= \max_{T \subseteq \mathcal{X}} \pi(T) - \mu(T)$$

$$T = \{y : \pi(y) > \mu(y)\}$$

$\rightarrow T \subseteq \mathcal{X}$

= "area b/w μ & π "

Properties:

$$0 \leq d_{TV}(\mu, \pi) \leq 1$$

$$d_{TV}(\mu, \pi) \leq d_{TV}(\mu, \nu) + d_{TV}(\nu, \pi)$$

Also written as: $\|\mu - \pi\|_{TV}$ or $d_{TV}(\mu, \pi)$

& if \mathcal{X} is infinite or continuous space,
then replace max over $S \subseteq \mathcal{X}$ by sup.

(2)

How to bound variation distance?

Coupling method.

For μ & π on \mathcal{X} ,

distribution ω on product space $\mathcal{X} \times \mathcal{X}$ is a coupling of μ & π if:

(rows sum to μ) - for all $i \in \mathcal{X}$, $\sum_{j \in \mathcal{X}} \omega(i, j) = \mu(i)$

(columns sum to π) - for all $j \in \mathcal{X}$, $\sum_{i \in \mathcal{X}} \omega(i, j) = \pi(j)$

Example: $\mathcal{X} = \{1, 2, 3, 4\}$

$$\mu = \left(\frac{1}{2}, \frac{1}{4}, 0, \frac{1}{4}\right) \quad \pi = \left(\frac{1}{3}, \frac{1}{3}, \frac{1}{3}, 0\right)$$

$$d_{TV}(\mu, \pi) = \frac{1}{2} \left(\frac{1}{6} + \frac{1}{12} + \frac{1}{3} + \frac{1}{4}\right) = \frac{5}{12}$$

$$\omega = \begin{bmatrix} \frac{1}{3} & \frac{1}{12} & \frac{1}{12} & 0 \\ 0 & \frac{1}{8} & \frac{1}{8} & 0 \\ 0 & 0 & 0 & 0 \\ 0 & \frac{1}{8} & \frac{1}{8} & 0 \end{bmatrix} \begin{matrix} = \frac{1}{2} \\ = \frac{1}{4} \\ = 0 \\ = \frac{1}{4} \end{matrix}$$

" $\frac{1}{3}$ " $\frac{1}{3}$ " $\frac{1}{3}$ "0"

$$\omega' = \begin{bmatrix} \frac{1}{3} & \frac{1}{12} & \frac{1}{12} & 0 \\ 0 & \frac{1}{4} & 0 & 0 \\ 0 & 0 & 0 & 0 \\ 0 & 0 & \frac{1}{4} & 0 \end{bmatrix} \begin{matrix} = \frac{1}{2} \\ = \frac{1}{4} \\ = 0 \\ = \frac{1}{4} \end{matrix}$$

" $\frac{1}{3}$ " $\frac{1}{3}$ " $\frac{1}{3}$ "

Both ω & ω' are couplings of μ, π .

(3)

For coupling ω of μ, π , take

sample $(\sigma, \tau) \sim \omega$

then $\sigma \sim \mu$ & $\tau \sim \pi$

(but they can be correlated with each other)

if uncorrelated (independent) then product coupling:

$$\omega'' = \begin{bmatrix} \frac{1}{6} & \frac{1}{6} & \frac{1}{6} & 0 \\ \frac{1}{12} & \frac{1}{12} & \frac{1}{12} & 0 \\ 0 & 0 & 0 & 0 \\ \frac{1}{12} & \frac{1}{12} & \frac{1}{12} & 0 \\ \frac{1}{3} & \frac{1}{3} & \frac{1}{3} & 0 \end{bmatrix} = \begin{matrix} \frac{1}{2} \\ \frac{1}{4} \\ 0 \\ \frac{1}{4} \end{matrix}$$

Coupling Lemma: For coupling ω of μ, π ,

let $(\sigma, \tau) \sim \omega$.

Then:

a) $d_{TV}(\mu, \pi) \leq \Pr(\sigma \neq \tau)$

b) \exists coupling ω where $d_{TV}(\mu, \pi) = \Pr(\sigma = \tau)$.

In words:

a) Any coupling upper bounds the variation distance.

b) There is always an optimal coupling.

In the previous example, note

$$\text{for } \omega: \Pr_{\omega}(\sigma \neq \tau) = \left(\frac{1}{12} + \frac{1}{12} + \frac{1}{8} + \frac{1}{8} + \frac{1}{8}\right) = \frac{13}{24} \geq \frac{5}{12}$$

$$\text{for } \omega': \Pr_{\omega'}(\sigma \neq \tau) = \left(\frac{1}{12} \times 2 + \frac{1}{4}\right) = \frac{5}{12} = d_{TV}(\mu, \pi)$$

So ω' is optimal.

Proof of (a): for $x \in \mathcal{X}$, $\omega(x, x) \leq \mu(x)$ & $\omega(x, x) \leq \pi(x)$

$$\text{thus, } \omega(x, x) \leq \min\{\mu(x), \pi(x)\}$$

$$\text{Thus, } \Pr(\sigma = \tau) = \sum_{\forall x} \omega(x, x) \leq \sum_x \min\{\mu(x), \pi(x)\}$$

$$\begin{aligned} \text{So, } \Pr(\sigma \neq \tau) &= 1 - \Pr(\sigma = \tau) \geq 1 - \sum_x \min\{\mu(x), \pi(x)\} \\ &= \sum_x \mu(x) - \min\{\mu(x), \pi(x)\} \\ &= \max_{s \in \mathcal{X}} \mu(s) - \pi(s) = d_{TV}(\mu, \pi). \quad \square \end{aligned}$$

(5)

Proof of (b): we'll construct an optimal ω :

Set $\omega(x, x) = \min\{u(x), \pi(x)\}$

for off-diagonal entries set ω to be a product of remaining mass.

Consider a MC defined by P on \mathcal{Z} .
Two chains (X_t) & (Y_t) .

Define joint evolution $(X_t, Y_t) \rightarrow (X_{t+1}, Y_{t+1})$

where $\Pr(X_{t+1}=j | X_t=i) = P(i, j)$

& $\Pr(Y_{t+1}=l | Y_t=k) = P(k, l)$

& if $X_t = Y_t$ then $X_{t+1} = Y_{t+1}$.

By the coupling lemma:

$$d_{TV}(P^+(X_0, \cdot), P^+(Y_0, \cdot)) \leq \Pr(X_+ \neq Y_+)$$

Let

$$T_{\text{couple}} = \max_{i, j \in \mathcal{Z}} \min \left\{ t : \Pr(X_+ \neq Y_+ | X_0 = i, Y_0 = j) \leq \frac{1}{4} \right\}$$

Then, $T_{\text{mix}} \leq T_{\text{couple}}$.

RW on n -dimensional hypercube $\{0, 1\}^n$. Let $\mathcal{Z} = \{0, 1\}^n$

From $X_+ \in \mathcal{Z}$,

1. Pick $i \in \{1, \dots, n\}$ var & $b \in \{0, 1\}$ var.

2. Set $X_{t+1}(i) = b$ & $X_{t+1}(j) = X_t(j)$
for $j \neq i$.

Consider 2 copies of this MC: (X_t) & (Y_t) . ⑦

Product coupling:

if $X_t = Y_t$: choose same i, b .

else choose i, b for X_t

& independently choose i', b' for Y_t .

Let $A_t = \{j : X_t(j) = Y_t(j)\}$ = "agree bits"

$D_t = \{j : X_t(j) \neq Y_t(j)\}$ = "disagree"

~~Note D_{t+1}~~

if $i \in D_t$ or $i' \in D_t$ then with Prob. $\frac{1}{2}$ agree afterwards

but if $i \in A_t$ or $i' \in A_t$ then w.p. $\frac{1}{2}$ disagree after.

So if $|D_t| = 1$ then Prob. $\approx \frac{1}{2n}$ that $X_{t+1} = Y_{t+1}$

& Prob. $\approx \frac{1}{2} |D_{t+1}| \geq 2$

Roughly:

$\frac{1}{n}$ $\frac{1}{2}$

this takes exponential time to reach 0.

~~1 1 1~~
0 1 2 3 ... n D_t

Better coupling: (identity coupling)

- choose same i, b for X_t, Y_t

if $i \in D_t$ then $|D_{t+1}| = |D_t| - 1$

& if $i \in A_t$ then $|D_{t+1}| = |D_t|$

Note, $D_{t+1} \subseteq D_t$.

$$E[|D_{t+1}| | X_t, Y_t] = |D_t| - \frac{|D_t|}{n} = |D_t| \left(1 - \frac{1}{n}\right)$$

$$\Pr(X_t \neq Y_t) \leq E[|D_t|] \quad \text{because if } X_t \neq Y_t \text{ then } E[|D_t|] \geq 1$$

$$\leq n \left(1 - \frac{1}{n}\right)^t$$

$$\leq n \left(1 - \frac{1}{n}\right)^t$$

$$\leq n e^{-t/n}$$

$$\leq \frac{1}{4} \text{ for } t \geq n \ln(4n).$$

$$\sum_p T_{mix} = O(n \log n).$$

Top-to-Random shuffle:

(9)

$\Omega =$ all $n!$ permutations of $\{1, \dots, n\}$

RW: take top card & insert into
random position i .

Ergodic: aperiodic & irreversible

but not symmetric so why is

$\pi = \text{uniform}(\Omega)$?

Note, P is doubly stochastic

(n transitions out of $X \in \Omega$

& n transitions into $X \in \Omega$

& every transition has prob. $1/n$).

HW: π is uniform iff P is doubly stochastic.

(see last class notes for

Proof in case of top-to-random
shuffle)

Simpler chain: Random-to-top

choose random card c & move to top.

Coupling: X_t & Y_t choose same c .

After choose c then c is on top for both,

& c is always in same spot for both.

So just need to choose ~~each~~ ^{every} card ≥ 1 time.

let $T =$ time to get every card at least once.

$t_i =$ time to get ~~the~~ i^{th} new card after collecting $(i-1)^{\text{st}}$.

then $T = \frac{1}{n} t_1 + t_2 + \dots + t_n$

for t_i , need to choose any of $n-(i-1)$ cards

(11)

$$\text{So let } p_i = 1 - \frac{i-1}{n} = \frac{n-i+1}{n}$$

$\& t_i = \text{geometric}(p_i)$

$$E[t_i] = \frac{1}{p_i} = \frac{n}{n-i+1}$$

$$\begin{aligned} E[T] &= \sum_{j=1}^n E[t_j] = \frac{n}{n} + \frac{n}{n-1} + \dots + \frac{n}{1} \\ &= n \left(1 + \frac{1}{2} + \frac{1}{3} + \dots + \frac{1}{n} \right) \\ &\leq n(1 + \ln n). \end{aligned}$$

Markov's inequality: $\Pr(T > 4E[T]) \leq \frac{1}{4}$

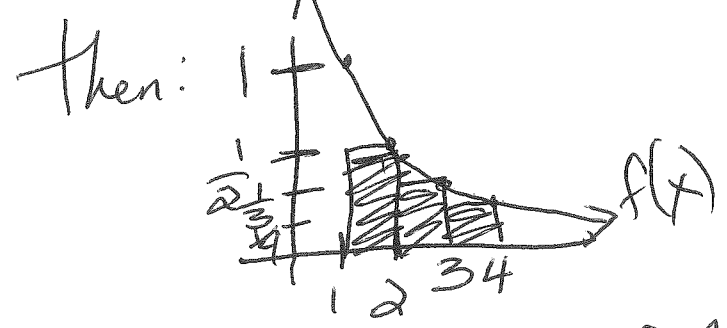
$$\begin{aligned} \text{So, } T_{\text{mix}} &\leq 4n(1 + \ln n) = 4n \ln(en) \\ &= O(n \log n). \end{aligned}$$

□

Claim: $1 + \frac{1}{2} + \dots + \frac{1}{n} \leq 1 + \ln n$

Proof: suffices to show: $\frac{1}{2} + \frac{1}{3} + \dots + \frac{1}{n} \leq \ln n.$

let $f(x) = \frac{1}{x}$



$$\frac{1}{2} + \dots + \frac{1}{n} = \sum_{j=2}^n \frac{1}{j} \leq \int_{x=1}^n \frac{1}{x} dx = \ln x \Big|_{x=1}^n = \ln n \quad \square$$