

CS292F StatRL Lecture 12

OPE in Bandits and Reinforcement Learning

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Spring 2021

UC Santa Barbara

Recap: Lecture 11

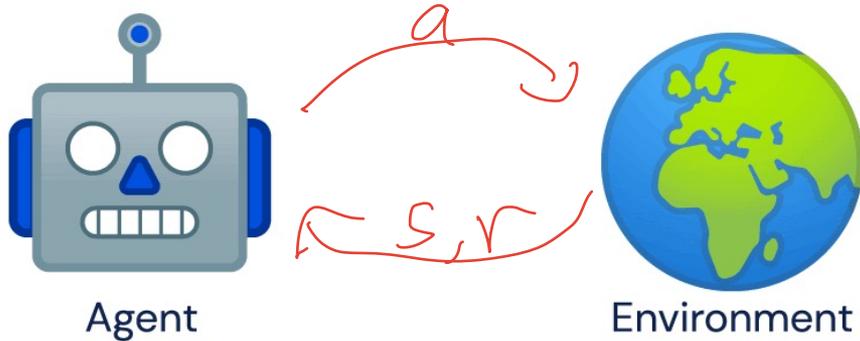
- Exploration in Linear MDPs
 - Finished the regret analysis
 - Uniform convergence with a covering number argument
- Short introduction to offline RL

What I did not cover about exploration?

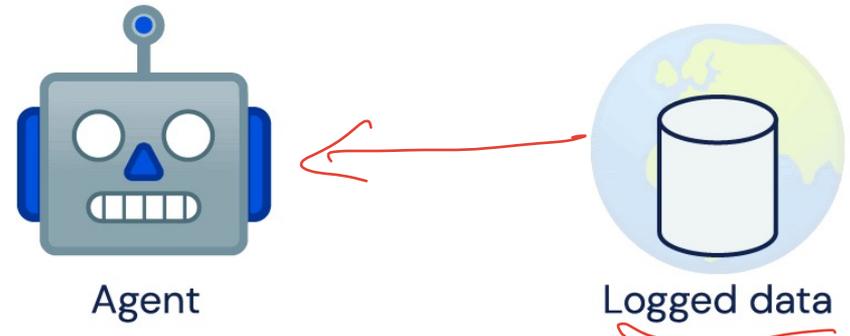
- Theoretical driven techniques
 - We covered: Optimism / UCB / Exploration bonuses
 - In the homework, you will see something else
 - We did not see: Thompson sampling
- Exploration techniques used in Deep RL
 - Curiosity (and other ways to add bonuses)
 - Adding noise (to model parameters / values / actions)
 - Random Network Distillation (user model-uncertainty)

Recap: Online RL vs Offline RL

Online Reinforcement Learning



Offline Reinforcement Learning

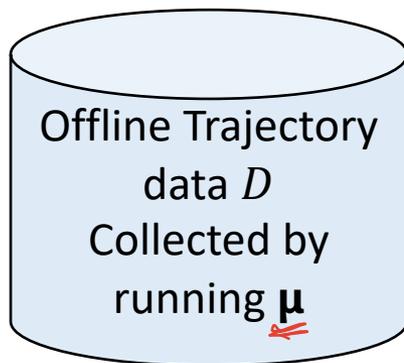


Exploration is often **expensive**, **unsafe**, **unethical** or **illegal** in practice, e.g., in self-driving cars, or in medical applications.

Can we learn a policy from already **logged interaction data**?

Recap: Offline Reinforcement Learning, aka. Batch RL

- Task 1: Offline Policy Evaluation. (OPE)

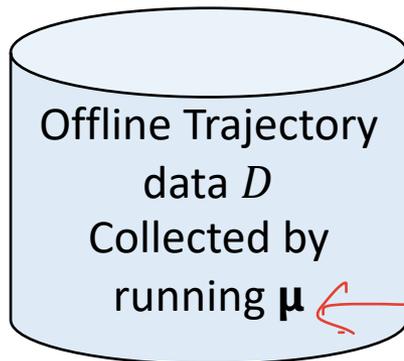


Task: design OPE methods

Evaluate fixed Target Policy π

Via Uniform OPE

- Task 2: Offline Policy Learning. (OPL)



Task: design OPO methods

Find near optimal Policy $\hat{\pi}^*$

logging policy

Recap: Contextual bandits model

- Contexts:

- $x_1, \dots, x_n \sim \lambda$ drawn iid, possibly infinite domain

- Actions:

- $a_i \sim \mu(a|x_i)$ Taken by a randomized "Logging" policy

- Reward:

- $r_i \sim D(r|x_i, a_i)$ Revealed only for the action taken

- Value:

- $v^\mu = \mathbb{E}_{x \sim \lambda} \mathbb{E}_{a \sim \mu(\cdot|x)} \mathbb{E}_D [r|x, a]$

- We collect data $(x_i, a_i, r_i)_{i=1}^n$ by the above processes.

$$(S_i, A_i, R_i)$$

Recap: Two standard approaches

$\hat{r}: S \times A \rightarrow \mathbb{R}$

- Direct method / regression estimator

$\hat{r}(x_i, a) \approx r(x_i, a)$

$$\hat{v}_{\text{DM}}^{\pi} = \frac{1}{n} \sum_{i=1}^n \sum_{a \in A} \hat{r}(x_i, a) \pi(a|x_i)$$

E_{π} plug in

- Importance sampling / Inverse Propensity Score

$E_{\pi} [R_i | x_i]$

$$\hat{v}_{\text{IPS}}^{\pi} = \frac{1}{n} \sum_{i=1}^n \frac{\pi(a_i|x_i)}{\mu(a_i|x_i)} r_i$$

$p_i(a_i|x_i)$

This lecture

- Continue with the estimators for OPEs in Bandits
 - Ideas to improve upon IS / DM
 - Some statistical analysis / comparisons
- OPE estimators for RL

Analyzing the performance of importance sampling estimator

$$\hat{V}_{IS}^{\pi} = \frac{1}{n} \sum_{i=1}^n \frac{\pi(A_i|S_i)}{\mu(A_i|S_i)} R_i$$

- Mean:

$$E[\hat{V}_{IS}^{\pi}] = \frac{1}{n} \sum_{i=1}^n E\left[\frac{\pi(A_i|S_i)}{\mu(A_i|S_i)} R_i\right] = E_{S_1, a} \left[E_{A_i, \mu(\cdot|S_i)} \left[\frac{E[R_i|S_i, A_i]}{r(S_i, A_i)} \frac{\pi(A_i|S_i)}{\mu(A_i|S_i)} \right] \right]$$

- Variance:

$$\begin{aligned} \text{Var}\left[\frac{1}{n} \sum_i \frac{\pi(A_i|S_i)}{\mu(A_i|S_i)} R_i\right] &= \frac{1}{n} \text{Var}\left[\frac{\pi(A_1|S_1)}{\mu(A_1|S_1)} R_1\right] = V^{\pi} \\ &= \frac{1}{n} \left(E_S \left[\text{Var}_{A_i, \mu(\cdot|S_i)} \left[\frac{\pi(A_i|S_i)}{\mu(A_i|S_i)} R_i \mid S_i \right] \right] + \text{Var}_S \left[E_{A_i, \mu(\cdot|S_i)} \left[\frac{\pi(A_i|S_i)}{\mu(A_i|S_i)} R_i \mid S_i \right] \right] \right) \\ &= \frac{1}{n} \left(E_S \left[E_{A_i, \mu(\cdot|S_i)} \left[\rho^2 \text{Var}[R_i|S_i, A_i] \mid S_i \right] + \text{Var}_{A_i, \mu(\cdot|S_i)} \left[E[\rho R_i|S_i, A_i] \mid S_i \right] \right] + \text{Var}_S \left[E_{\pi} [r(S_i, A_i) \mid S_i] \right] \right) \\ &= \frac{1}{n} \left(\underbrace{E_{\mu} [\rho^2 G^2(S, A)]}_{\text{Reward Variance}} + \underbrace{E \left[\text{Var}_{\mu} [\rho r(S, A_i)] \mid S \right]}_{\text{Log-policy's Variance}} + \underbrace{\text{Var} \left[E_{\pi} [r(S, A_i) \mid S_i] \right]}_{\text{Variance of the context}} \right) \end{aligned}$$

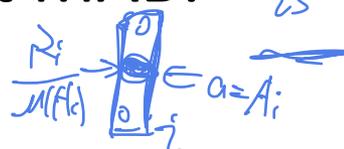
Consider an even simpler setting:
Multi-armed bandits with K-arms

$$\hat{v}_{\text{DM}}^{\pi} = \frac{1}{n} \sum_{i=1}^n \sum_{a \in \mathcal{A}} \hat{r}(x_i, a) \pi(a|x_i) = \sum_{a \in \mathcal{A}} \bar{r}(a) \bar{\pi}(a)$$

$$\hat{v}_{\text{IPS}}^{\pi} = \frac{1}{n} \sum_{i=1}^n \frac{\pi(a_i|x_i)}{\mu(a_i|x_i)} r_i$$

Importance Sampling and Direct Method are surprisingly similar in some cases

- Consider just MAB:



$$r_{ES}(a) = \begin{cases} 0 & \text{if } A_i \neq a \\ \frac{R_i}{\mu(A_i)} & \text{if } A_i = a \end{cases}$$

$$f_{IS}(a) = \begin{cases} \frac{1}{n} \sum_i \frac{R_i \mathbb{1}(A_i=a)}{\mu(A_i)} \\ = \frac{1}{n} \sum_i \hat{r}_{IS}(a) \end{cases}$$

- Importance sampling as a regression estimator

$$\hat{V}_{IS}^{\pi} = \frac{1}{n} \sum_i \frac{\pi(A_i)}{\mu(A_i)} R_i = \sum_{a \in \mathcal{A}} \pi(a) \underbrace{\frac{1}{n} \sum_i \frac{R_i \mathbb{1}(A_i=a)}{\mu(A_i)}}_{f_{IS}(a)} = \sum_{a \in \mathcal{A}} \pi(a) \hat{r}_{IS}(a)$$

- Regression estimator as an importance sampling

$$\hat{V}_{DM}^{\pi} = \sum_a \hat{r}(a) \cdot \pi(a) = \sum_a \frac{1}{N_a} \sum_i R_i \mathbb{1}(A_i=a) \cdot \pi(a) = \sum_a \frac{1}{n} \frac{1}{\mu(a)} \sum_i R_i \mathbb{1}(A_i=a) \cdot \pi(a)$$

$$= \frac{1}{n} \sum_i \frac{\pi(a)}{\mu(a)} \cdot R_i \mathbb{1}(A_i=a)$$

$$\hat{r}(a) = \begin{cases} \frac{1}{N_a} \sum_i R_i \mathbb{1}(A_i=a) & \text{if } N_a > 0 \\ 0 & \text{otherwise} \end{cases}$$

- Which one is better?

$$N_a = \frac{\sum \mathbb{1}(A_i=a)}{n} = n \cdot \mu(a)$$

$$E[R|A=a] = r(a)$$

$$\hat{r}(a) \xrightarrow{N_a \rightarrow \infty} r(a)$$

Comparing the MSE of DM and IS

In Simple Math: IS

$$MSE(\hat{V}_{IS}) = \text{Var}[\hat{V}_{IS}] = \frac{1}{n} \left[\underbrace{E_{\mu(A)} \left[\frac{\sigma(A)^2}{\mu(A)} \right]}_{V_1} + \underbrace{\text{Var}_{\mu(A)} \left[\frac{\mu(A)}{\mu(A)} \cdot r(A) \right]}_{V_2} \right]$$

- Mean Square Error and Bias-Variance

Decomposition

$$MSE(\hat{V}) = E[(\hat{V} - V^*)^2] = E[(\hat{V} - E[\hat{V}] + E[\hat{V}] - V^*)^2]$$

$$= E[(\hat{V} - E[\hat{V}])^2] + E[(E[\hat{V}] - V^*)^2] + 2E[(\hat{V} - E[\hat{V}])(E[\hat{V}] - V^*)]$$

\downarrow
 $\text{Var}(\hat{V})$ $(\text{Bias}(\hat{V}))^2$ $(E[\hat{V} - V^*])$

- Analyzing DM with plug-in estimator

$$\hat{V}_{DM} = \sum_a \pi(a) \hat{r}(a)$$

$$E[\hat{V}_{DM}] = \sum_a \pi(a) E[\hat{r}(a)] = \sum_a \pi(a) E\left[\frac{1}{N_a} \sum R_i \mathbb{1}(A_i=a) \mathbb{1}(N_a > 0)\right]$$

$$= \sum_a \pi(a) P(N_a > 0) E\left[\frac{1}{N_a} \sum R_i \mathbb{1}(A_i=a) \mid N_a > 0\right]$$

$$\text{Bias}(\hat{V}_{DM}) = E[\hat{V}_{DM}] - V^* = \sum_a \pi(a) r(a) [P(N_a > 0) - 1] \xrightarrow{n \rightarrow \infty} 0$$

$$\text{Var}\left[\sum_a \pi(a) \hat{r}(a)\right] = E\left[\text{Var}\left[\sum_a \pi(a) \hat{r}(a) \mid N_a, a \in A\right]\right] + \text{Var}\left[E\left[\sum_a \pi(a) \hat{r}(a) \mid N_a, a \in A\right]\right]$$

$$= E\left[\sum_a \frac{\pi(a)^2}{N_a} \text{Var}(\hat{r}(a) \mid N_a)\right] + \text{Var}\left[\sum_a \pi(a) r(a) \mathbb{1}(N_a > 0)\right]$$

$$\approx E\left[\sum_a \frac{\pi(a)^2}{N_a} \sigma^2(a) \mathbb{1}(N_a > 0)\right] + R_{\max}^2 \cdot P(\exists a, N_a > 0)$$

$$\leq E\left[\sum_a \frac{\pi(a)^2}{n \cdot \pi(a)} \sigma^2(a) \mathbb{1}(N_a > 0)\right] + R_{\max}^2 \cdot P(\exists a, N_a > 0)$$

$$= \frac{1}{n} \sum_a \pi(a) \sigma^2(a) + R_{\max}^2 \cdot P(\exists a, N_a > 0)$$

$$\frac{\sum_a \pi(a)^2 \sigma^2(a)}{n \pi(a)} + \frac{1}{n} \sum_a \pi(a) \sigma^2(a)$$

Weighted importance sampling

- Self-normalization

$$\hat{V}_{wIS} = \frac{1}{\sum_{i=1}^n P_i} \sum_{i=1}^n P_i R_i \quad P_i = \frac{\pi(A_i|S_i)}{\mu(A_i|S_i)}$$

$$E\left[\sum_{i=1}^n P_i\right] = E\left[\sum_{i=1}^n \frac{\pi(A_i|S_i)}{\mu(A_i|S_i)}\right] = \sum_{i=2}^n \sum_{\alpha} \mu(A_i|S_i) \frac{\pi(A_i|S_i)}{\mu(A_i|S_i)} = n$$

$$\hat{V}_{wIS} \xrightarrow{n \rightarrow \infty} \sqrt{n}$$

Experiment 1: Facebook data

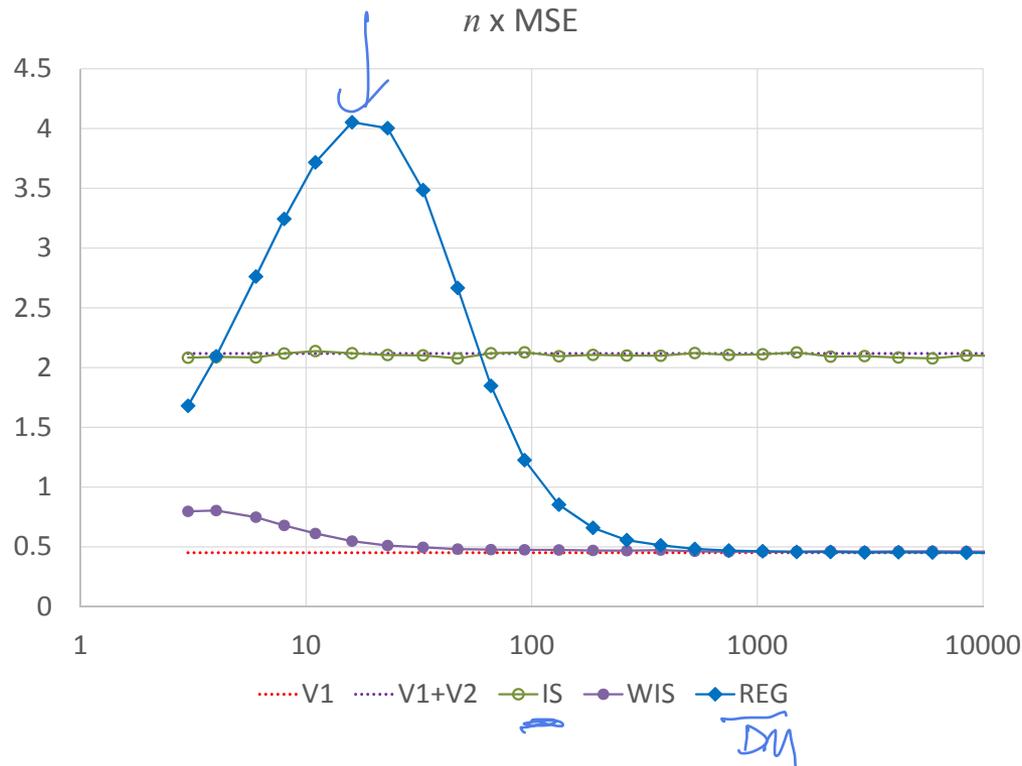


Figure 3: $n\text{MSE}$ for query “facebook” ($K = 2178$). The asymptotic rates V_1 and $V_1 + V_2$ are provided for reference.

Experiment 2: Gmail data

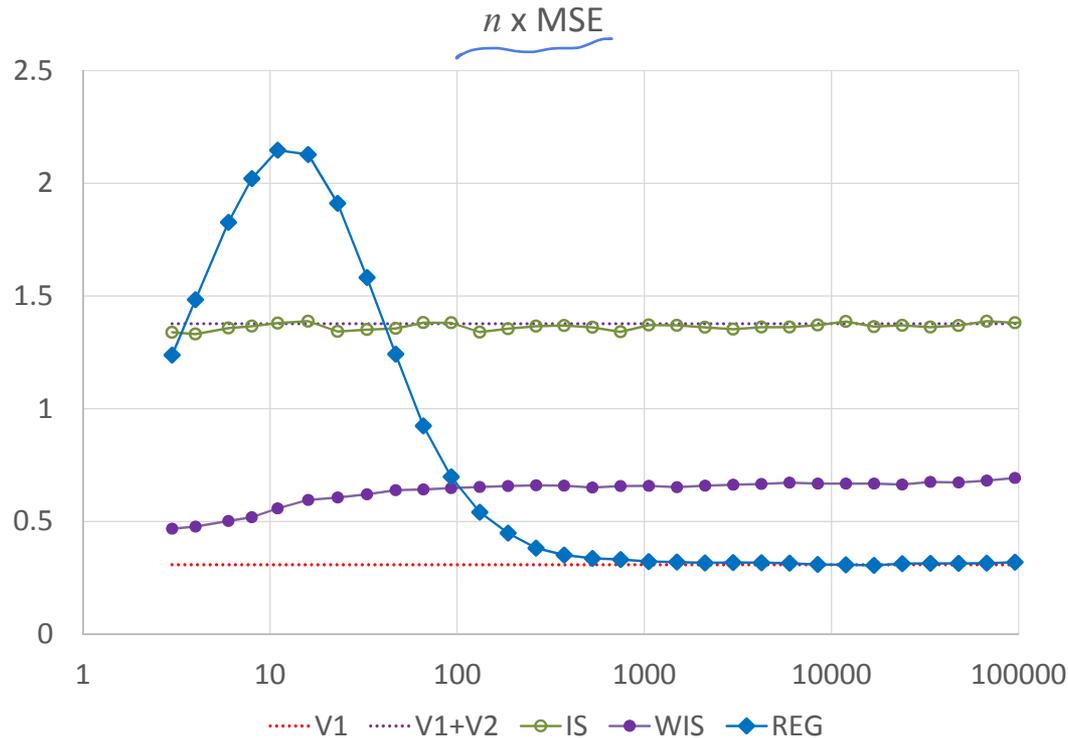


Figure 4: $n\text{MSE}$ for query “gmail” ($K = 648$). The asymptotic rates V_1 and $V_1 + V_2$ are provided for reference.

(x_i, a_i, v_i)

Doubly robust estimator for OPE

(Contextual bandits)

(S_i, A_i, R_i)

- We are using the regression estimator as a baseline.

$$\hat{V}_{DR} = \frac{1}{n} \sum_{i=1}^n \left(\underbrace{E_{\pi_i}[\hat{V}(S_i, A_i) | S_i]}_{\text{baseline}} + \frac{P_i}{\pi_i} (R_i - \hat{V}(S_i, A_i)) \right)$$

$\hat{V}(S, a)$ $V(S, a)$

fixed / given

estimated from a different data

$$\begin{aligned} E[\hat{V}_{DR}] &= E_{S_i \sim \pi} \left[E_{A_i}[\hat{V}(S_i, A_i) | S_i] - E \left[\sum_a \pi(a|S_i) \cdot \frac{\pi(a|S_i)}{\pi(a|S_i)} \cdot E[R_i | S_i, a] \right] \right] \\ &= E[V(S_i, A_i)] = V^* \end{aligned}$$

$\sum_a \pi(a|S_i) \hat{V}(S_i, a)$

Theory of doubly robust estimator

unknown $\mu_i = \mu(A_i|S_i) \in \hat{\mu}(S_i; A_i)$

unknown $r(S_i) \rightarrow r(S_i)$
 $h(S_i) \rightarrow \mu(S_i)$

- Double robustness in model-misspecification

$$\hat{V}_{DR} = \frac{1}{n} \sum_i \hat{r}^{\pi}(S_i) + \frac{\pi(A_i|S_i)}{\hat{\mu}(S_i; A_i)} \cdot (R_i - \hat{r}(S_i; A_i))$$

if either $\hat{\mu}$ or \hat{r} is consistent, then $\hat{V}_{DR} \rightarrow V^{\pi}$

If either $E[\hat{\mu}] = \mu$ or $E[\hat{r}] = r$, then $E[\hat{V}_{DR}] = V^{\pi}$

$\hat{r}(S_i; A_i) \rightarrow r(S_i; A_i)$
 \parallel
 $E[R_i | S_i; A_i]$

- Variance reduction (sometimes)

\hat{V}_{DR} is asymptotically efficient

if ~~$(\hat{r} - r) \cdot (\frac{1}{\hat{\mu}} - \frac{1}{\mu})$~~

$$(\hat{r} - r) \cdot \left(\frac{1}{\hat{\mu}} - \frac{1}{\mu}\right) \Rightarrow 0 \left(\frac{1}{\sqrt{n}}\right)$$

$\hat{r} = o_p(1)$ if $\hat{\mu} = \mu$

(Robins and Rotnitzky, 1995; Bang and Robins, 2005)

Lower bounding the minimax risk

- Our main theorem: assume λ is a probability density, then under mild moment conditions

$$\inf_{\hat{v}} \sup_{D(r|a,x) \in \mathcal{R}(\sigma^2, R_{\max})} \mathbb{E}(\hat{v} - v^\pi)^2$$
$$= \Omega \left[\underbrace{\frac{1}{n} \left(\mathbb{E}_\mu[\rho^2 \sigma^2] \right)}_{\text{Randomness in reward}} + \underbrace{\mathbb{E}_\mu[\rho^2 R_{\max}^2]}_{\text{Randomness due to context distribution}} \right]$$

S is optimal in the worst case

Classical optimality theory (Hahn, 1998)

- n^* Var[any LAN estimator] is greater than:

$$\mathbb{E}_{x \sim \mathcal{D}} \{ \mathbb{E}_{\mu} [\rho^2 \text{Var}(r|x, a)|x] \} + \text{Var}_{x \sim \mathcal{D}} \{ \mathbb{E}_{\mu} [\rho r|x] \} .$$

Take  supremum

$$\underline{\mathbb{E}_{\mu} [\rho^2 \sigma^2] + \mathbb{E}_{x \sim \mathcal{D}} [\mathbb{E}_{\mu} [\rho R_{\max}|x]^2]} .$$

- The minimax lower bound is bigger!

$$\mathbb{E}_{\mu} [\rho^2 \sigma^2] + \mathbb{E}_{\mu} [\rho^2 R_{\max}^2]$$

How could that be? There are estimators that achieve asymptotic efficiency.

- e.g., [Robins](#), [Hahn](#), [Hirano](#), [Imbens](#), and many others in the semiparametric efficiency industry!

Assumption:	Realizable assumption: $E[r x,a]$ is differentiable in x for each a .	No assumption on $E[r x,a]$ except boundedness.
Consequences	Hirano et. al. is optimal. Imbens et. al. is optimal. IPS is suboptimal!	IPS is optimal (up to a universal constant)
Caveat	Poor finite sample performance. Exponential dependence in d .	Does NOT adapt to easier problems.

SWITCH estimator

- Recall that IPS is bad because: $\hat{v}_{\text{IPS}}^{\pi} = \frac{1}{n} \sum_{i=1}^n \frac{\pi(a_i|x_i)}{\mu(a_i|x_i)} r_i$

- SWITCH estimator:

For each $i = 1, \dots, n$, for each action $a \in \mathcal{A}$

if $\pi(a|x_i)/\mu(a|x_i) \leq \tau$:

Use IPS (or DR).

else:

Use regression estimator.

Handwritten notes in blue ink:

- Left: $r_i = \begin{bmatrix} r_i \\ 0 \\ \vdots \\ 0 \end{bmatrix} \in A_i$
- Middle: A_i
- Right: $r_i(s_i, a) = \begin{bmatrix} r_i(s_i, a) \\ \vdots \\ r_i(s_i, a) \end{bmatrix}$
- Bottom: $r_i = \begin{bmatrix} r_i(s_i, a) \\ \vdots \\ r_i \end{bmatrix}$

Error bounds for SWITCH

$$\text{MSE}(\hat{v}_{\text{SWITCH}}) \leq$$

$$\frac{2}{n} \mathbb{E}_{\mu} \left[\underbrace{(\sigma^2 + R_{\max}^2) \rho^2 \mathbf{1}(\rho \leq \tau)}_{(1)} \right]$$

$$+ \frac{2}{n} \mathbb{E}_{\pi} \left[\underbrace{R_{\max}^2 \mathbf{1}(\rho > \tau)}_{(2)} \right]$$

$$+ \underbrace{\mathbb{E}_{\pi} \left[\epsilon \mathbf{1}(\rho > \tau) \right]^2}_{(3)}$$

1) Variance from IPS (reduced by truncation)

2) Variance due to sampling x . Required even with perfect oracle

1) Bias from the oracle.

Automatic parameter tuning

- Conservative approximate MSE minimizing.

$$\hat{\tau} = \underset{\tau}{\operatorname{argmin}} \widehat{\operatorname{Var}}_{\tau} + \widehat{\operatorname{Bias}}_{\tau}^2.$$

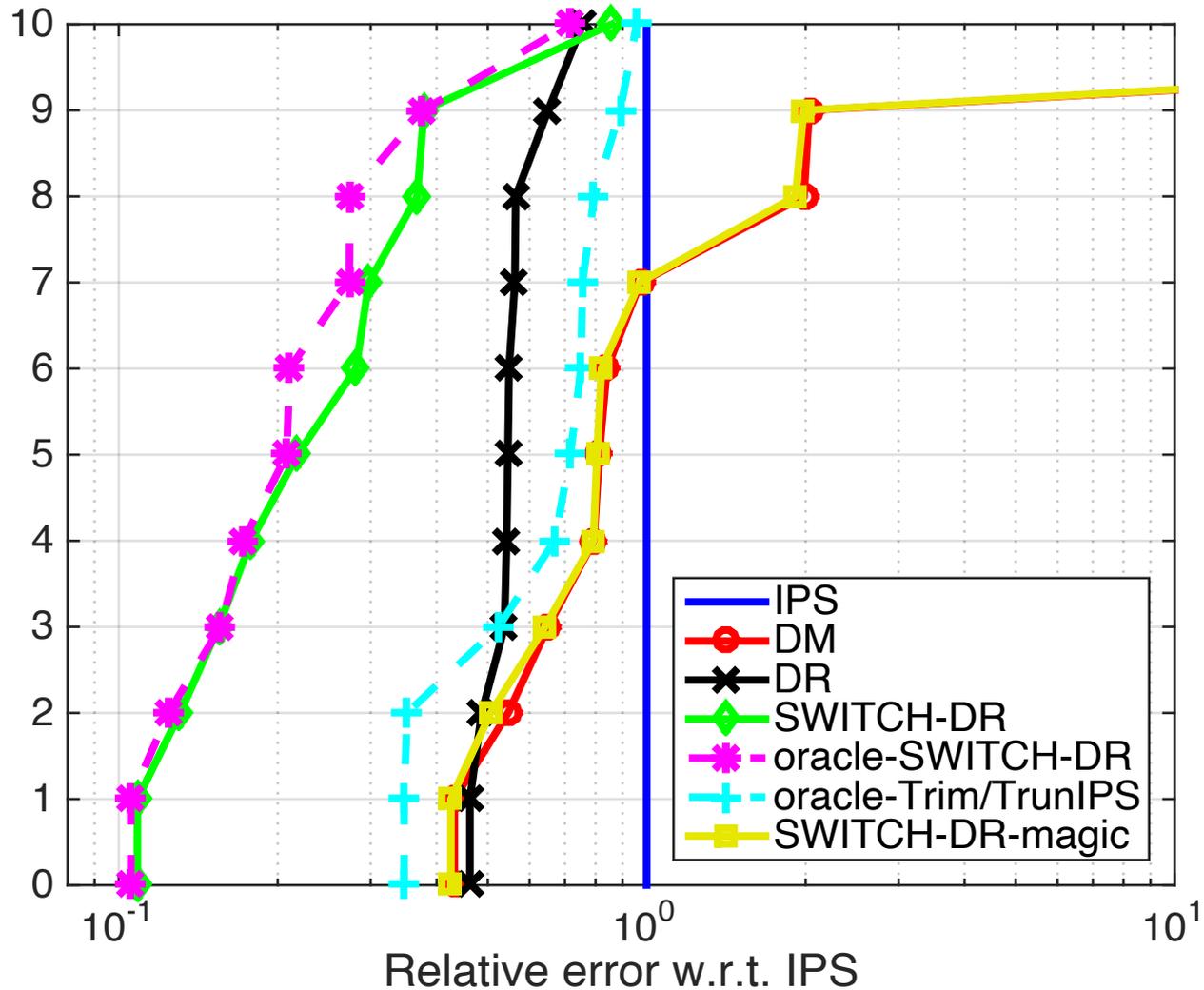
- Details:

$$Y_i(\tau) := r_i \rho_i \mathbf{1}(\rho_i \leq \tau) + \sum_{a \in \mathcal{A}} \hat{r}(x_i, a) \pi(a|x_i) \mathbf{1}(\rho(x_i, a) > \tau) \quad \text{and} \quad \bar{Y}(\tau) = \frac{1}{n} \sum_{i=1}^n Y_i(\tau),$$

$$\operatorname{Var}(\hat{v}_{\text{SWITCH}-\tau}) = \frac{1}{n} \operatorname{Var}(\hat{v}_{\text{SWITCH}-\tau}(x_1)) \approx \frac{1}{n^2} \sum_{i=1}^n (Y_i(\tau) - \bar{Y}(\tau))^2 =: \widehat{\operatorname{Var}}_{\tau},$$

$$\begin{aligned} \operatorname{Bias}^2(\hat{v}_{\text{SWITCH}}) &\leq \mathbb{E}_{\mu}[\rho \epsilon^2 | \rho > \tau] \pi(\rho > \tau)^2 \leq \mathbb{E}_{\mu}[\rho R_{\max}^2 | \rho > \tau] \pi(\rho > \tau)^2 \\ &\approx \left[\frac{1}{n} \sum_{i=1}^n \mathbb{E}_{\pi}(R_{\max}^2 | \rho > \tau, x_i) \right] \left[\frac{1}{n} \sum_{i=1}^n \pi(\rho > \tau | x_i) \right]^2 =: \widehat{\operatorname{Bias}}_{\tau}^2. \end{aligned}$$

With additional label noise



Checkpoint: OPE for Contextual Bandits

- Estimators: DM, IS, WIS, DR, SWITCH
- Bias-Variance Tradeoff
- Optimality theory
 - Depends on whether you have access to a good regression estimator